

Exam D0M61A Advanced econometrics

19 January 2009, 9–12am

Question 1 (5 pts.)

Consider the wage function

$$w_i = \beta_0 + \beta_1 S_i + \beta_2 E_i + \beta_3' h_i + \varepsilon_i,$$

where w_i is the log-wage of individual i , S_i is years of schooling, E_i is years of experience, h_i is a vector of other observed characteristics of i , and ε_i is an error term. Using labour market data from 1976 from the U.S. National Longitudinal Survey of Young Men (Card, 1995), the following equations were estimated by OLS, with standard errors in parentheses (2040 observations were used):

$$w_i = 4.8159 + 0.0825S_i + 0.0436E_i + \text{residual} \quad (1)$$

(0.0824) (0.0046) (0.0028)

$$R^2 = 0.1450 \quad s = 0.3864 \quad SSR = 304.256$$

$$w_i = 4.8920 + 0.0732S_i + 0.0427E_i$$

(0.0813) (0.0045) (0.0027)

$$- 0.171BLACK_i + 0.155SMSA_i - 0.086SOUTH_i + \text{residual} \quad (2)$$

(0.024) (0.019) (0.018)

$$R^2 = 0.2093 \quad s = 0.3719 \quad SSR = 281.359$$

Here $BLACK_i$, $SMSA_i$ and $SOUTH_i$ are dummy variables indicating whether the individual was black, lived in a metropolitan area and lived in the south.

1. ($\frac{1}{2}$ pt.) Comment on the following assertion: “From a prediction perspective, where one wants to predict someone’s wage given his level of schooling, experience and other observed characteristics, there is no point in worrying about endogeneity.”

2. ($\frac{1}{2}$ pt.) Give a precise interpretation of the coefficient estimate associated with S_i in (1).
3. (1 pt.) Assuming that the conditions for exact inference are satisfied, test the joint hypothesis that the coefficients associated with $BLACK_i$, $SMSA_i$ and $SOUTH_i$ are zero. (For hypotheses tests, always formally state the null hypothesis, the alternative hypothesis, the test statistic, the chosen significance level, the critical value and the outcome. Choose the significance level to be as informative as possible or, even better, compute the p -value.)
4. (1 pt.) Comment on the following assertion: “The above estimates of β_1 are not helpful for an individual who wants to estimate his *returns to schooling*, even when all individuals have the *same* returns to schooling.” Start your comments with a sensible definition of the returns to schooling for any individual i (in terms of i ’s wage only!).

Motivated by a desire to estimate the returns to schooling, equation (2) was re-estimated with GMM, using the following instruments: a constant, $BLACK_i$, $SMSA_i$, $SOUTH_i$, and also AGE_i (age), KWW_i (score on Knowledge of the World of Work test), IQ_i (score on IQ test) and $NEARCOL_i$ (dummy variable indicating whether i lived near a college in 1966). The results are as follows (with heteroskedasticity-robust standard errors in parentheses):

$$\begin{aligned}
 w_i = & 4.4504 + 0.1054S_i + 0.0425E_i \\
 & (0.1239) \quad (0.0081) \quad (0.0027) \\
 & - 0.1372BLACK_i + 0.1405SMSA_i - 0.0834SOUTH_i + \text{residual} \quad (3) \\
 & (0.0260) \quad (0.0196) \quad (0.0187) \\
 R^2 = & 0.1792 \quad s = 0.3789 \quad SSR = 292.069
 \end{aligned}$$

5. ($\frac{1}{2}$ pt.) The estimate of the returns to schooling has increased. Which critical assumption related to the effect of ability do we have to make for this estimate to be reliable? What is the likely sign of the bias if this assumption fails to be true?

6. ($\frac{1}{2}$ pt.) Why is it that, next to S_i , also E_i is considered to be endogenous here?

7. ($\frac{1}{2}$ pt.) What is the motivation behind the use of $NEARCOL_i$ as an instrument?

8. ($\frac{1}{2}$ pt.) The J -statistic equals 7.958. Test the hypothesis that the instruments are jointly valid.

Question 2 (3 pts.)

Cameron and Trivedi (*Microeconometrics: Methods and Applications*, Cambridge University Press, 2005) write on p. 68:

“If this conditional mean is linear in x , so that $E[y|x] = x'\beta$, the parameter β has a structural or causal interpretation [...] This permits meaningful policy analysis of effects of changes in regressors on the conditional mean.”

Comment on this statement.

Question 3 (6 pts.)

Consider a random sample of n workers, drawn from the population of workers who become unemployed in a given country, in a given period (say, January 2008). Let y_i be the length of worker i 's unemployment spell, i.e. the time elapsed before i finds a new job. We seek to explain y_i by i 's characteristics, denoted as x_i (a vector). Assume $y_i \geq 0$ is a random variable whose conditional density, given x_i , is

$$f(y_i|x_i; \lambda_i) = \lambda_i \exp(-\lambda_i y_i), \quad \text{where } \lambda_i = \exp(\beta_0 + \beta'x_i)$$

and β_0 and β (a vector) are parameters to estimate.

1. (1 point) Suppose that one of the variables in x_i is the amount of education i has taken (expressed in years, say). What is the likely sign of the parameter (say β_j) associated with this variable? Motivate your answer.

2. (2 points) Given the observations $(y_1, x_1), \dots, (y_n, x_n)$, construct the log-likelihood function of (β_0, β) .

3. (2 points) Consider the following complication. Suppose that for some workers in the sample, the observed unemployment spell is incomplete due to the fact that these workers have not found a new job by the time the observation period ends (say, 31 December 2008). To account for this, let y_i be i 's complete (but possibly unobserved) unemployment spell, and let z_i be the observed, possibly incomplete, unemployment spell. Here z_i is related to y_i by

$$\begin{aligned} z_i &= y_i && \text{if } i\text{'s unemployment spell is complete} \\ z_i &< y_i && \text{if } i\text{'s unemployment spell is incomplete} \end{aligned}$$

and we know whether or not z_i is a complete or an incomplete unemployment spell. That is, we observe the variable d_i defined as

$$\begin{aligned} d_i &= 0 && \text{if } i\text{'s unemployment spell is complete} \\ d_i &= 1 && \text{if } i\text{'s unemployment spell is incomplete.} \end{aligned}$$

Given the observations $(z_1, x_1, d_1), \dots, (z_n, x_n, d_n)$, construct the log-likelihood function of (β_0, β) .

4. (1 point) Suppose $n = 1$ and suppose there are no covariates, so we only observe (z_i, d_i) for a single individual i . If $d_i = 1$, what is the maximum likelihood estimate of i 's expected unemployment duration? [You may want to reflect on the following equivalent question. Suppose you execute a command on your computer, without the slightest idea of how much time it will take before execution stops. A natural model for the waiting time before execution stops is the exponential distribution (recall it is memoryless). Suppose you wait for 30 seconds and execution hasn't stopped. What is the maximum likelihood estimate of the expected waiting time? What is it after only waiting for 0.5 seconds?]

Question 4 (6 pts.)

- (3 points) Suppose you wish to estimate the population correlation between X and Y , defined as

$$r = \frac{E[(X - E(X))(Y - E(Y))]}{\sqrt{E(X - E(X))^2 E(Y - E(Y))^2}}.$$

Suppose you only have 10 observations, $(X_1, Y_1), \dots, (X_{10}, Y_{10})$. It is well known that the sample correlation,

$$\hat{r} = \frac{\sum_{i=1}^{10} (X_i - \bar{X})(Y_i - \bar{Y})}{\sqrt{\sum_{i=1}^{10} (X_i - \bar{X})^2 \sum_{i=1}^{10} (Y_i - \bar{Y})^2}},$$

is a biased estimator of r . Can you use the bootstrap to reduce the bias of \hat{r} ? Describe your procedure in detail. How do you solve the (potential) problem that a sample correlation is only defined when the denominator is non-zero?

2. (2 points) Describe how you obtain a bootstrap 95% confidence interval for the IV estimator $\hat{\delta} = (X'Z)^{-1}X'y$ in the just-identified case (using the notation of the course notes).

3. (1 point) Is it possible to apply the parametric bootstrap to the Hansen's J test? If so, describe the procedure briefly. If not, indicate why.